



## When attention is intact in adults with ADHD

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### Abstract

Is covert visuospatial attention—selective processing of information in the absence of eye movements—preserved in adults with attention-deficit/hyperactivity disorder (ADHD)? Previous findings are inconclusive due to inconsistent terminology and sub-optimal methodology. To settle this question, we used well-established spatial cueing protocols to investigate the perceptual effects of voluntary and involuntary attention on an orientation discrimination task for a group of adults with ADHD and their neurotypical age-matched and gender-matched controls. In both groups, voluntary attention significantly improved accuracy and decreased reaction times at the relevant location, but impaired accuracy and slowed reaction times at irrelevant locations, relative to a distributed attention condition. Likewise, involuntary attention improved accuracy and speeded responses. Critically, the magnitudes of all these orienting and reorienting attention effects were indistinguishable between groups. Thus, these counter-intuitive findings indicate that spatial covert attention remains functionally intact in adults with ADHD.

**Keywords** ADHD · Visual perception · Selective attention · Orienting · Reorienting · Microsaccades · Performance fields

Despite symptoms of “inattention” as a qualitatively defining feature (American Psychiatric Association, 2013), surprisingly, much is still unknown about attention in attention-deficit/hyperactivity disorder (ADHD). ADHD is a neurodevelopmental disorder characterized by a heterogeneous set of persistent maladaptive behaviors and neurocognitive impairments. Initially conceptualized as a behavioral disorder of hyperactivity and heightened impulsivity in children, the notion of deficient

attention was introduced in the third edition of the *Diagnostic and Statistical Manual of Mental Disorders* (American Psychiatric Association, 1980; Barkley, 2007). Many years later, it was recognized that ADHD may persist through adolescence, and even onset in adulthood (Barkley, 2007). ADHD is estimated to affect 5% to 6% of the worldwide adult population (Polanczyk, de Lima, Horta, Biederman, & Rohde, 2007), and a burgeoning body of literature shows that adults with ADHD exhibit abnormalities in several domains, including, but not limited to, response precision, cognitive flexibility, working memory, temporal information processing, response inhibition, and our cognitive process of interest, attention (Hervey, Epstein, & Curry, 2004; Mueller, Hong, Shepard, & Moore, 2017; Pievsky & McGrath, 2017).

Attention is not a unitary concept (Carrasco, 2011; Posner, 2014). The cognitive requirements of many of the “classic” tasks adopted to probe attentional functioning in ADHD (e.g., continuous performance task, Stroop tasks, Attention Network Test) involve several distinct attentional and executive functions (Hervey et al., 2004; Mueller et al., 2017). For example, high performance on Stroop and flanker tasks, in which observers are asked to ignore distracting features or other stimuli, requires strong executive interference control and response inhibition in addition to intact selective attention. This combination of task demands is unfortunate, as patients’ most reliable deficits lie under the umbrella of executive functions, which encompass response inhibition, reward response, decision-making, and motivational processes, among others

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(Barkley, 1997; Willcutt, Doyle, Nigg, Faraone, & Pennington, 2005; Pievsky & McGrath, 2017). Thus, any of these factors could be responsible for observed behavioral differences between groups. Moreover, terminology spanning clinical and experimental adult ADHD research has been inconsistent and imprecise.

There are significant gaps regarding which types of attention are deficient in ADHD. Research in adults with ADHD on sustained attention reports significant limitations in their ability to continuously perform a task over a prolonged period (e.g., minutes; Dankner, Shalev, Carrasco, & Yuval-Greenberg, 2017; Marchetta, Hurks, De Sonneville, Krabbendam, & Jolles, 2007; Mueller et al., 2017). There has been surprisingly little research on selective attention (i.e., the preferential processing of one stimulus in the presence of other distracting stimuli; Mueller et al., 2017; Tsal, Shalev, & Mevorach, 2005). A few studies have argued that ADHD adults show heightened distractibility to irrelevant distractors (Godefroid & Wiersema, 2017; Marchetta et al., 2007; Marzinzik et al., 2012; Tucha et al., 2008). In contrast to the rich body of literature on visuospatial orienting in children (e.g., Ortega, López, Carrasco, Anllo-Vento, & Aboitiz, 2013; Tsal et al., 2005; see a review by Huang-Pollock & Nigg, 2003), only five studies have explicitly investigated the perceptual effects of spatial “orienting” and “reorienting” of selective attention (Posner, 1980) in adults with ADHD (Dhar, Been, Minderaa, & Althaus, 2008; Epstein, Conners, Erhardt, March, & Swanson, 1997; Epstein, Johnson, Varia, & Conners, 2001; Oberlin, Alford, & Marrocco, 2005; Tomporowski, Tinsley, & Hager, 1994). These studies suggest that orienting and reorienting may be functionally spared in the disorder, but suboptimal methodology complicates their interpretation (see General Discussion).

For the first time, we isolate, manipulate and measure the visual perceptual effects of both covert voluntary (endogenous) and involuntary (exogenous) spatial attention in adults with mild-to-moderate ADHD and their age-matched and gender-matched neurotypical (NT) controls. The perceptual consequences of these types of attention are often the same in NT observers; both increase contrast sensitivity, enhance spatial resolution, accelerate the rate of information accrual, and even alter stimulus appearance. However, these two types of selective attention can differ according to task demands and stimuli, exhibit different temporal dynamics, and are supported by partially overlapping and interactive yet distinct neural networks (Carrasco, 2011). Given that these types of attention can result in different perceptual consequences (Carrasco, 2011) and that their functional roles differ, it is critical to assess both types of attention in this population as only one, both, or none could be spared. Understanding whether these types of attention are fully functional or impaired in the ADHD brain is informative for developing more sophisticated neuropsychological and neurobiological models

of the disorder. We measured both accuracy and RT for an orientation discrimination task. In addition, we compared microsaccades (MS) of these two groups as they differ in continuous sustained tasks (Dankner et al., 2017; Fried et al., 2014) and have been linked to some perceptual and attentional tasks (Rucci & Poletti, 2015). In two separate psychophysical experiments, we used central and peripheral cues to directly manipulate either endogenous (Experiment 1) or exogenous (Experiment 2) attention.

## Experiment 1: Endogenous attention

### Method

**Observers** To be included in the study, all adult observers had to possess normal or corrected-to-normal vision. Observers in the ADHD group were clinically diagnosed with DSM-IV-TR ADHD (American Psychiatric Association, 2000) according to the Adult ADHD Clinical Diagnostic Scale (Adler & Spencer, 2004), and Structured Clinical Interview for DSM-IV, Research Version, Non-patient Edition (SCID-I/NP; First, Spitzer, Gibbon, & Williams, 2002). We did not exclude potential observers on the basis of age, race, gender, ADHD severity, or comorbidities. Fourteen adults with ADHD (see Table S1 in the Supplementary Information) and 14 age-matched and gender-matched NT controls ( $M_{\text{age}} = 31.0$  years,  $SD = 8.5$ ; seven female) participated in Experiment 1; all observers had attended college and some graduate school. Our sample size was similar to previous studies reporting significant performance differences between ADHD and NT groups in purported tasks of focused and sustained attention (for reviews, see Frazier, Demaree, & Youngstrom, 2004; Schoechlin & Engel, 2005), and similar or larger than studies that found intact attention effects in other special populations (autism: Grubb, Behrmann, Egan, Minshew, Carrasco & Heeger, 2013a, b; amblyopia: Roberts, Cymerman, Smith, Kiorpes, & Carrasco, 2016) and neurotypical observers (e.g., Carrasco, Ling, & Read, 2004; Carrasco, Talgar, & Cameron, 2001; Carrasco & Yeshurun, 1998; Dugué, Roberts, & Carrasco, 2016; White, Rolfs, & Carrasco, 2015). All experimental procedures were in agreement with the Helsinki declaration and approved by the New York University and NYU School of Medicine Institutional Review Boards. All observers (except for author M.R., control observer) were naïve to the experimental hypotheses and signed written consent.

**Apparatus and setup** Observers were tested in the same dimly lit, sound-attenuated room for both experiments. Stimuli were programmed on an Apple iMac MC413LL/A 21.5-in. Desktop (3.06 GHz Intel Core 2 Duo) using MATLAB (MathWorks, Natick, Massachusetts, USA) in conjunction with the MGL toolbox (<http://justingardner.net/mgl>). They were presented at a viewing distance of 57 cm on a 21-in.

IBM P260 CRT monitor (1280 × 960 pixel resolution, 90-Hz refresh rate), calibrated and linearized using a Photo Research (Chatworth, CA) PR-650 SpectraScan Colorimeter. Observers performed the experiments using a forehead and chin rest to ensure head stabilization. Eye movements were monitored using an EyeLink 1000 Desktop Mount eye tracker (SR Research, Ontario, Canada).

**Stimuli** Observers were asked to fixate on a black, centrally placed cross (0.5° across) throughout the trial (see Fig. 1). Four placeholders—each composed of four black dots (0.05° radius) arranged in a circle 0.5° from the location of an upcoming Gabor patch stimulus (to prevent masking)—were always presented on the screen to reduce location uncertainty. The target and three distractor stimuli were all 3.2° wide, 4-cycles-per-degree Gabor patches (contrast-defined sinusoidal gratings embedded in a Gaussian envelope,  $\sigma = 0.46^\circ$ ), randomly and independently tilted  $\pm 20^\circ$  from vertical, centered at 6.4° eccentricity along the diagonals, and with the same mean luminance as the uniform gray background. To manipulate endogenous spatial attention, we presented a central precue—either a single 0.88° line or four 0.28° lines (all 0.14° thick)—0.38° from the center of the fixation cross, which pointed to one or all (neutral, distributed condition) of the possible target locations. The response cue indicated the target location by pointing to one placeholder (that matched the single central precue for valid trials and mismatched for invalid trials) and eliminated location uncertainty at the response time for all conditions.

**Procedure** Observers performed the same experimental procedure across two hour-long behavioral sessions. They completed about 18 blocks of 60 trials each for a total of 1,080 trials; 648 trials in the valid cue condition (60% of all trials), and 216 trials each in both the invalid (20% of all trials) and neutral cue (20% of all trials) conditions. At the beginning of the first session, observers completed practice blocks (24 trials each, 100% stimulus contrast) until they could perform the task reliably above chance. Then, they underwent a staircase procedure (neutral cues only) where we obtained their individual stimulus contrast thresholds yielding 80% accuracy. The contrast of the Gabor patch stimuli was initially set at each individual's threshold performance around 80%. The required stimulus contrast did not differ between the ADHD ( $M = 32.1 \pm 9.1\%$ ) and NT controls ( $M = 19.6 \pm 7.5\%$ ),  $t(26) = 1.07$ ,  $p > .1$ , Cohen's  $d = 0.4$ , scaled JZS Bayes factor = 1.9, according to Rouder, Speckman, Sun, Morey, and Iverson (2009). If observers made an eye movement  $\geq 1^\circ$  from the fixation cross between initiation and stimulus offset, the trial would immediately abort and the text, "Please fixate," would appear at the center of the screen. These trials were rerun at the end of the block. Both groups broke fixation (ADHD:  $M = 2.86 \pm 0.5$ ; NT:  $M = 2.40 \pm 0.9$ ) with similar frequency per block (independent-samples  $t$  test),  $t(26) = 0.45$ ,  $p > .1$ , Cohen's  $d = 0.2$ , scaled JZS Bayes factor = 2.6.

The ADHD group broke fixation in 4.8% of all trials; the NT group in 4.0%.

**Task and trial sequence** Observers performed a two-alternative forced-choice (2AFC) orientation discrimination task binocularly while endogenous spatial attention was manipulated via presentation of either a single (80% of all trials, of which 75% of trials were valid and 25% trials were invalid) or distributed central precue (20% of all trials; see Fig. 1). On every trial, observers were encouraged to respond as accurately as possible, without time stress. After 250 ms, the precue was presented for 400 ms, after which there was a brief inter-stimulus interval (ISI) of 60 ms. The 460-ms stimulus-onset-asynchrony (SOA) between precue onset and stimulus was designed to ensure that all observers had ample time to voluntarily deploy their endogenous attention (Liu, Stevens, & Carrasco, 2007; Müller & Rabbitt, 1989; Nakayama & Mackeben, 1989; see a review by Carrasco, 2011). After the interval, the target and distractor Gabor patches appeared simultaneously inside the placeholders for 120 ms.<sup>1</sup>

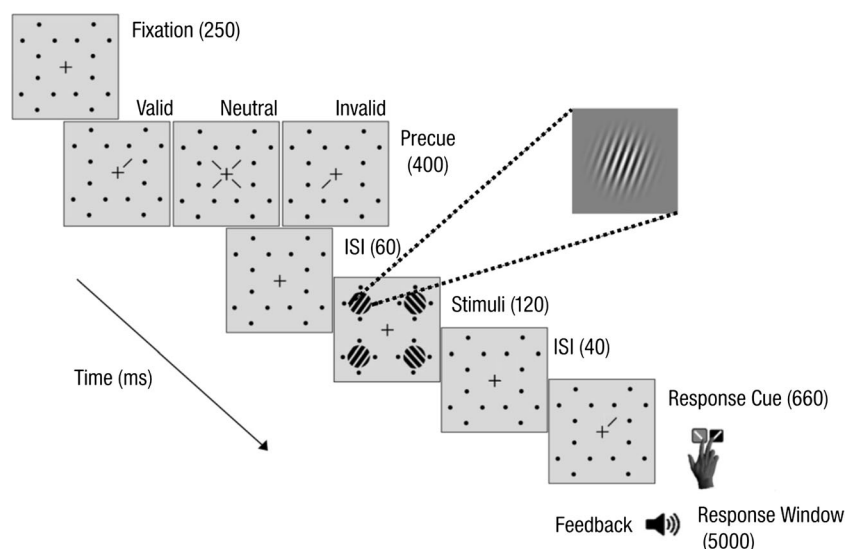
There was a brief 40-ms ISI between display offset and the response cue, which remained on the screen for 660 ms. An auditory tone indicated the beginning of the 5,000-ms response window, in which observers had to report the target orientation (clockwise or counterclockwise relative to vertical) using one of two keyboard presses (1 for clockwise, 2 for counterclockwise) with their right hand. Observer response terminated the response window, after which there was a mandatory 1,000-ms intertrial interval. Auditory feedback was provided at the end of each trial and visual feedback indicating observers' accuracy and number of fixation breaks was presented at the end of each block.

## Results

**Overall performance** Overall accuracy in the neutral cueing condition was similar in the ADHD ( $M = 79.3 \pm 2.1\%$ ) and NT ( $M = 79.3 \pm 1.6\%$ ) observers, confirming that task difficulty was matched across groups (see Fig 2a). A two-way mixed-design ANOVA revealed a main effect of cue,  $F(1.41, 36.8) = 41.3$ ,  $p < .001$ ,  $\eta_p^2 = .61$ , but neither the main effect of group,  $F(1, 26) = 1.1$ ,  $p > .1$ ,  $\eta_p^2 = .04$ , nor its interaction with cue condition ( $F < 1$ ,  $\eta_p^2 = .02$ ) were significant, indicating no differences in overall accuracy or the magnitude of the attention effect between groups.

To confirm that any nonsignificant results were not simply due to a lack of statistical power to find differences between groups, we calculated Bayesian information criterion probabilities (pBIC) that represent the strength of evidence in favor of the null (H0)—a nonsignificant main effect or interaction—or alternative (H1)—a significant main effect or interaction—

<sup>1</sup> Except for one 53-year-old observer, for whom the stimuli were displayed for 180 ms so that she could perform the task above chance.



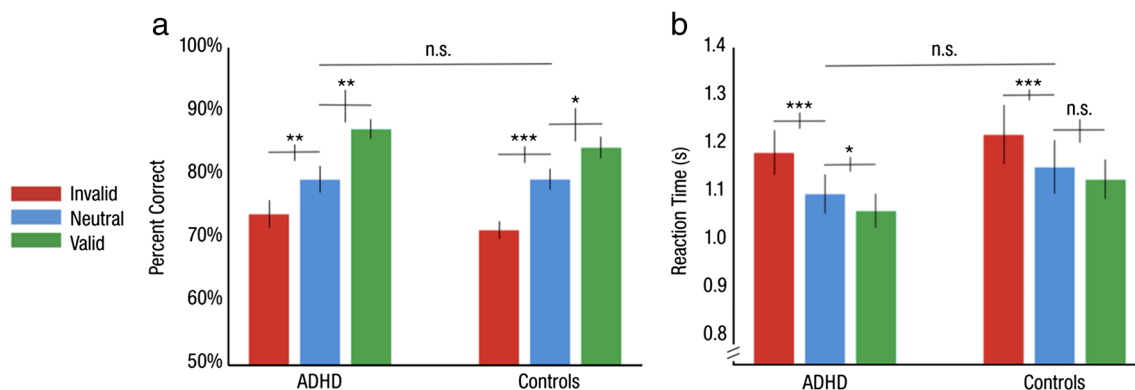
**Fig. 1** Trial sequence for Experiment 1: Endogenous attention

hypotheses given our data set D (Masson, 2011). A value of  $pBIC$  between .75 and .95 and a Bayes factor  $>3$  are considered positive evidence. The Bayes factor analysis of the main effect of group provided evidence in favor of the null with an odds of 1.55 to 1:  $pBIC(H0|D) = .61$  and  $pBIC(H1|D) = .39$ . The analysis of the interaction between cue and group, the test of greatest interest to this study, provided positive evidence in favor of the null hypothesis with an odds of 4.78 to 1:  $pBIC(H0|D) = .83$  and  $pBIC(H1|D) = .17$ .

A corresponding two-way mixed ANOVA of RT found a similar pattern of results (Fig 2b; see [Supplementary Information](#)). In both groups, accuracy was significantly higher for the valid (ADHD:  $M = 87.3 \pm 1.5\%$ ; NT:  $M = 84.4 \pm 1.6\%$ ) and lower for the invalid (ADHD:  $M = 73.9 \pm 2.1\%$ ; NT:  $M = 71.3 \pm 1.3\%$ ) compared to the neutral cueing condition (see Fig 2a). The Bayes factor analysis found positive evidence in support of the null hypothesis that there was no significant difference in overall RT between groups with an odds of 3.02 to 1:  $pBIC(H0|D) = .75$  and  $pBIC(H1|D) = .25$ . We also found evidence in favor of

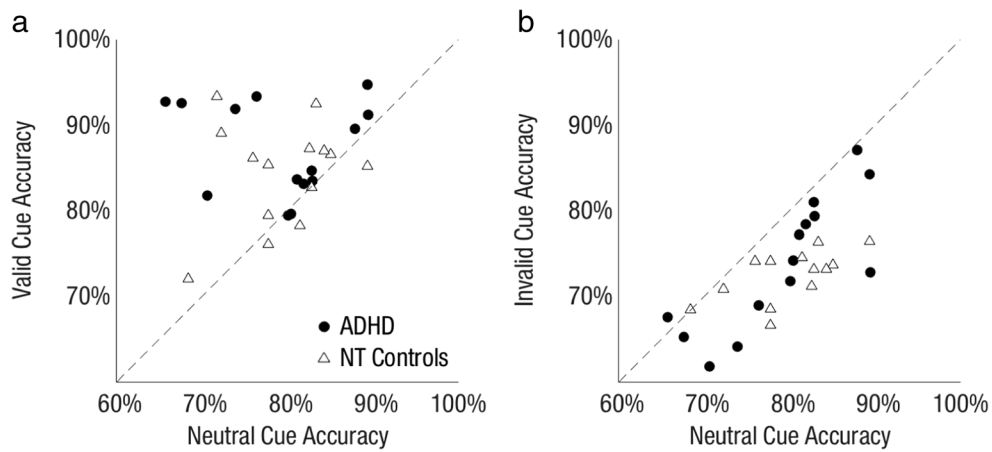
the alternative hypothesis of a significant Cue  $\times$  Group interaction, with the ADHD group exhibiting a greater benefit of the endogenous cue, faster responses during the valid than the neutral cueing condition ( $BF < .01$ ),  $pBIC(H0|D) < .01$ ,  $pBIC(H1|D) = .99$ . Figure 3 shows that the variance of individual endogenous attention benefits and costs were similar in both groups and present for all but a few observers, who fall on the diagonal line.

**Microsaccades** The average frequency of MS throughout the trial per block did not significantly differ among attention conditions (valid:  $M = 30.9 \pm 2.07$ , neutral:  $M = 30.8 \pm 2.29$  and invalid:  $M = 30.1 \pm 1.94$ ), or between groups (ADHD:  $M = 153.6 \pm 15.1$ , NT:  $M = 153.9 \pm 14.7$ ; both  $ps > .1$ ). MS in each condition followed the main sequence: the higher the amplitude, the faster the velocity (see Fig. 4, Column 1). Furthermore, their kinematics (i.e., peak velocity, amplitude, and duration) did not differ as a function of attention condition or group (all  $ps > .1$ ; see Fig. 4, Columns 2–4).

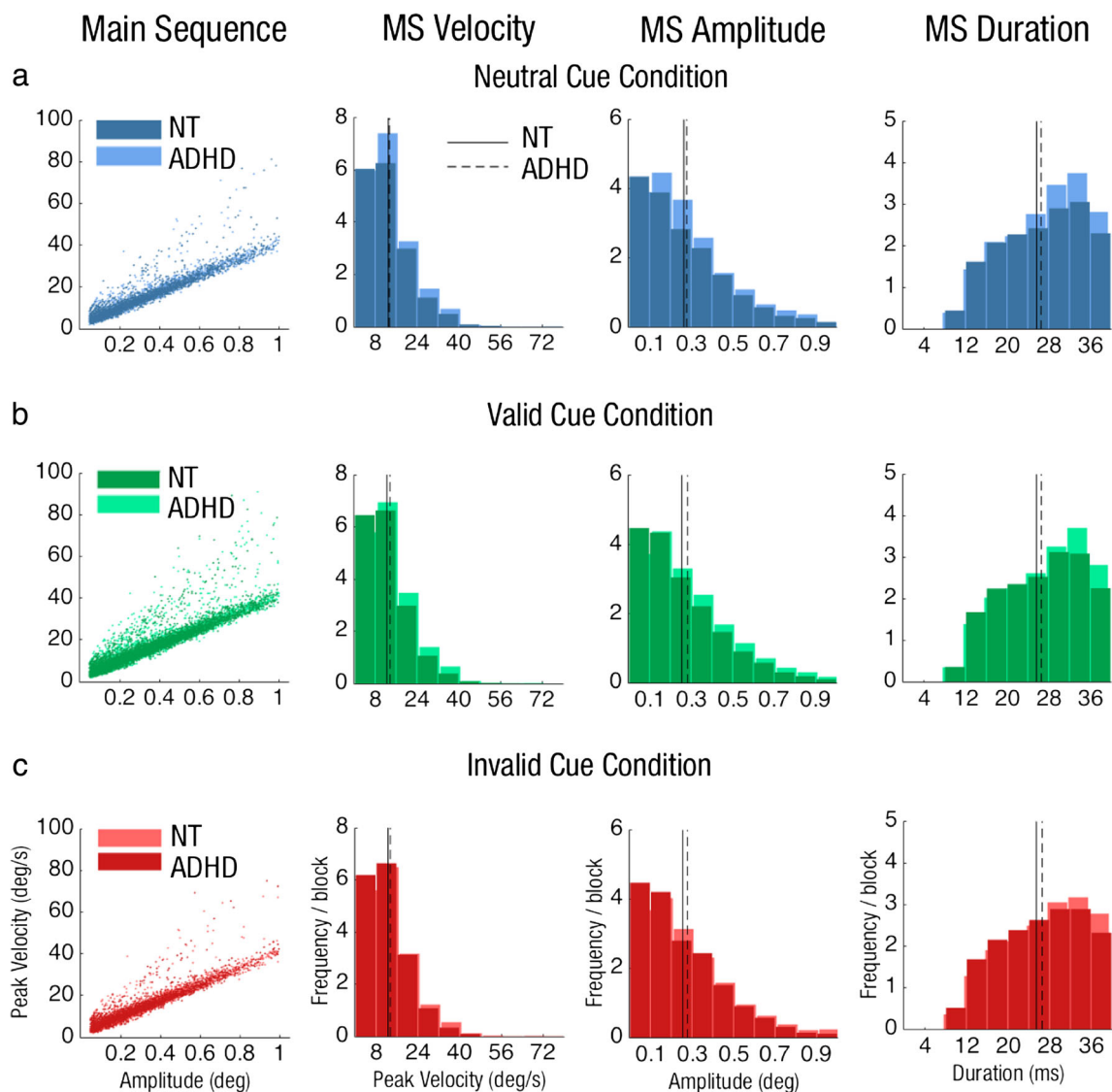


**Fig. 2** Performance in Experiment 1: Endogenous attention. a Percentage accuracy. b Reaction times. Error bars are  $\pm 1$  SEM. \* $p < .05$ . \*\* $p < .01$ . \*\*\* $p < .001$ . (Color figure online)





**Fig. 3** Performance accuracy for individual observers in Experiment 1. a Valid versus neutral cue condition. b Invalid versus neutral cue condition. The farther from the diagonal, the greater the attention (a) benefit and (b) cost



**Fig. 4** MS kinematics for both groups in Experiment 1: Endogenous attention for the neutral cue condition (a), valid cue condition (b), and invalid cue condition (c). MS frequency for the valid condition was divided by three to normalize the trial probability within each block.

Distributions of MS velocity (Column 2), amplitude (Column 3), and duration (Column 4) in terms of average frequency per block for each group. (Color figure online)

In sum, adults with ADHD demonstrate the classic benefit in both accuracy and RT of voluntarily orienting to a spatial location that they will be subsequently asked about, as well as the cost of initially orienting to the incorrect location. The magnitude and pattern of these typical attentional effects are indistinguishable from those found in NT adults. Moreover, the oculomotor correlates of these behavioral effects were similar between both groups.

## Experiment 2: Exogenous attention

In Experiment 1, we demonstrated that voluntary covert orienting and reorienting of selective visual attention remains functionally intact in adults with ADHD. Psychophysical, neuroimaging, and neurophysiological studies indicate that the two types of attention are supported by interactive and partially overlapping yet distinct neural systems (Carrasco, 2011; Posner, 2014). In this experiment, we investigated whether exogenous selective attention is also preserved in adults with ADHD, employing essentially the same task as in Experiment 1. Moreover, our task design enabled us to directly assess, for the first time, the spatial distribution of attention across the visual field in adults with ADHD. Some studies on neuropsychological disorders that are often comorbid with ADHD (e.g., autism spectrum disorder; see, e.g., Keehn, Müller, & Townsend, 2013) have reported differences in the spatial distribution of attention across the visual field, but others have found a similar distribution (Grubb, Behrmann, Egan, Minshew, Carrasco & Heeger, 2013a, b). We investigated potential perceptual and attentional asymmetries at isoecentric locations (see [Supplementary Information](#)).

## Method

**Observers** Inclusion and exclusion criteria were the same as in Experiment 1. Fourteen adults with ADHD (six also participated in Experiment 1; see Table S1 in the [Supplementary Information](#)) and 14 age-matched and gender-matched NT controls ( $M_{\text{age}} = 30.9$ ,  $SD = 8.0$ ; seven female) participated in Experiment 2; all observers had attended college and some graduate school. The ADHD group was mainly composed of individuals exhibiting mild-to-moderate ADHD symptomology, according to the Conners' Adult ADHD Rating Scale (CAARS; Conners, Erhardt, & Sparrow, 1999).

**Apparatus and setup** Apparatus and setup were identical to those in Experiment 1.

**Stimuli** Stimuli were identical to those in Experiment 1, except for the stimuli locations and form of the precue (see Fig. 5). We presented the placeholders and Gabor patches at the cardinal axes. To manipulate exogenous attention, the dots of either one (valid peripheral precue) or all four (neutral precue) placeholders grew in size (to  $0.16^\circ$  radius), and the color changed from black to white. The response cue (a  $0.8^\circ$  line

placed  $0.3^\circ$  from the central fixation cross) indicated the target location by pointing to one placeholder (matching the peripheral precue location in the valid condition).

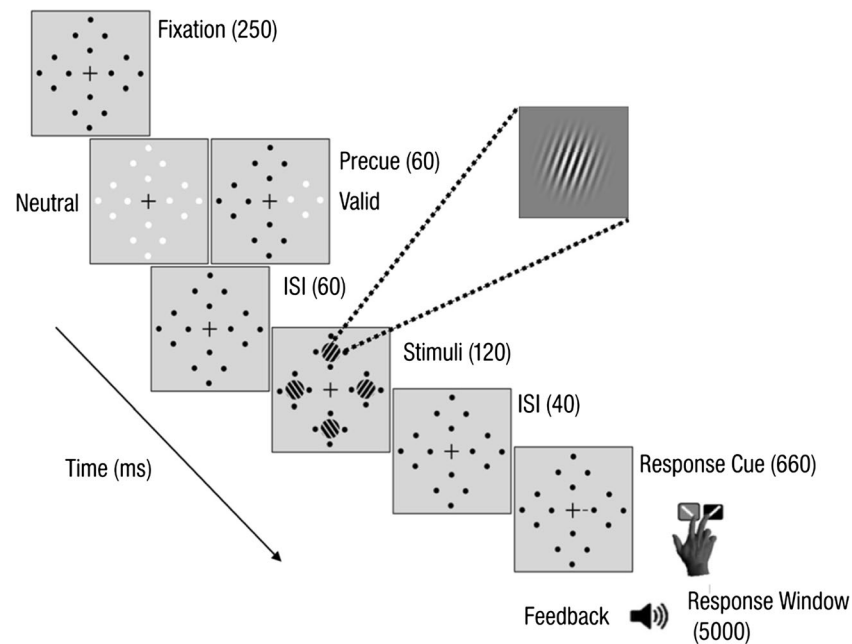
**Procedure** The procedure was identical to that of Experiment 1, except that (a) there were no invalid cues (with exogenous attention, the benefit of the valid cue is the same regardless of whether there are invalid trials or not; Carrasco, 2011; Giordano, McElree, & Carrasco, 2009) and (b) both groups of observers completed about 20 experimental blocks of 48 trials each. There was not a significant difference in the mean contrast required by the ADHD ( $M = 34.2 \pm 7.5\%$ ) and NT controls ( $M = 25.3 \pm 7.4\%$ ),  $t(26) = 0.8$ ,  $p > .1$ , scaled JZS Bayes factor = 2.2, according to Rouder et al. (2009). Trials in which observers broke fixation were cancelled and excluded from analyses. Both groups broke fixation (ADHD:  $M = 2.09 \pm 0.6$ ; NT:  $M = 1.02 \pm 0.4$ ), with similar frequency per block (independent-samples  $t$  test),  $t(26) = 1.45$ ,  $p > .1$ , scaled JZS Bayes factor = 1.3.

**Task and trial sequence** Observers performed the same 2AFC orientation discrimination task binocularly while exogenous spatial attention was manipulated via presentation of either a valid peripheral (50% of trials) or a neutral, distributed (50% of trials) precue (see Fig. 5). The sequence was the same as in Experiment 1, except that the precue duration was only 60 ms. The 120-ms SOA between precue onset and stimulus was designed to optimize the attentional effects of the exogenous cue and prevent any voluntary deployment of attention (Liu et al., 2007; Müller & Rabbitt, 1989; Nakayama & Mackeben, 1989; see a review by Carrasco, 2011).

## Results

**Overall performance** Once again, task difficulty in the neutral condition was well equated between groups (ADHD:  $M = 78.4 \pm 2.1\%$ ; NT:  $M = 81.5 \pm 1.3\%$ ; see Fig. 6a). A two-way mixed-design ANOVA of accuracy revealed a significant main effect of cue,  $F(1, 26) = 31.2$ ,  $p < .001$ ,  $\eta_p^2 = .55$ , but neither the main effect of group,  $F(1, 26) = 1.09$ ,  $p > .1$ ,  $\eta_p^2 = .04$ , nor their interaction ( $F < 1$ ,  $\eta_p^2 = .03$ ) was significant (see Fig. 6a). The Bayes factor analysis of the main effect of group provided evidence in favor of the null, with an odds of 2.22 to 1:  $\text{pBIC}(H_0|D) = .69$  and  $\text{pBIC}(H_1|D) = .31$ . The analysis of the interaction between cue and group, the test of greatest interest to this study, provided positive evidence in favor of the null hypothesis with an odds of 4.56 to 1:  $\text{pBIC}(H_0|D) = .82$  and  $\text{pBIC}(H_1|D) = .18$ .

A corresponding analysis of RT found a similar pattern of results (Fig. 6b; see [Supplementary Information](#)). In both groups, accuracy was higher in the valid (ADHD:  $M = 85.2 \pm 1.3\%$ ; NT:  $M = 86.4 \pm 1.8\%$ ) than in the neutral cueing condition (Fig. 6a), ADHD,  $t(13) = 4.25$ ,  $p = .001$ , 95% CI [3.4, 10.4], Cohen's  $d = 1.1$ ; NT,  $t(13) = 3.63$ ,  $p = .003$ , 95% CI [2.0, 7.8], Cohen's  $d = 1.0$ . The Bayes factor analysis found positive evidence in support of the null hypothesis that there



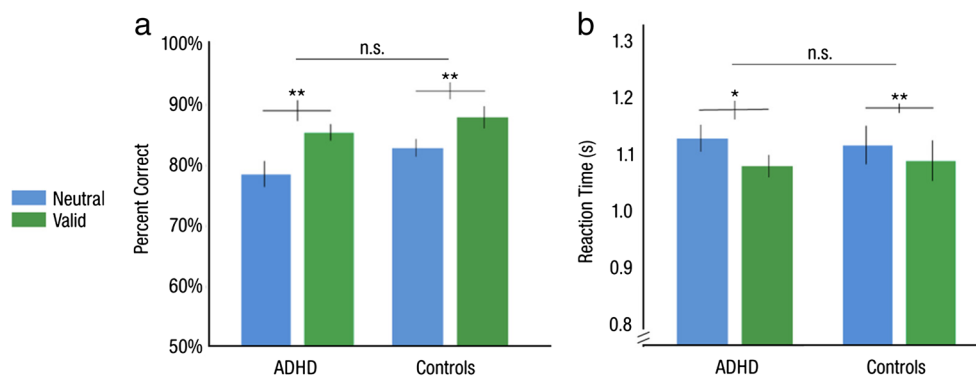
**Fig. 5** Trial sequence for Experiment 2: Exogenous attention

was no significant difference in overall RT between groups with an odds of 4.72 to 1:  $p\text{BIC}(H_0|D) = .83$  and  $p\text{BIC}(H_1|D) = .17$ . The Bayes factor analysis favored the null hypothesis regarding the interaction between cueing condition and group with odds of 2.60 to 1:  $p\text{BIC}(H_0|D) = .72$  and  $p\text{BIC}(H_1|D) = .28$ . Thus, both groups exhibited the classic exogenous attention benefit to the same extent. The variance of individual exogenous attention accuracy benefits was similar for both groups and present for all but a few observers, whose data are along the diagonal line (see Fig 7.).

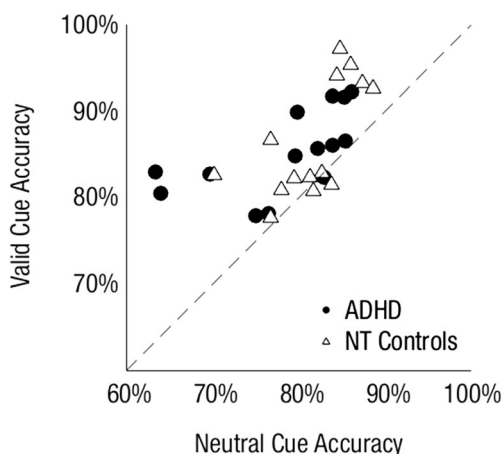
**Microsaccades** The average frequency of MS throughout the trial per block did not significantly differ between attention conditions (valid:  $M = 45.6 \pm 4.45$ , neutral:  $M = 43.6 \pm 4.35$ ;  $p > .1$ ) or between groups (ADHD:  $M = 73.5 \pm 9.8$ , NT:  $M = 105 \pm 13.6$ ;  $p > .07$ ). MS in all conditions followed the main

sequence; the higher the amplitude, the faster the speed (see Fig. 8, Column 1). Further, their kinematics (i.e., peak velocity, amplitude, and duration) did not differ as a function of attention condition or group (all  $p$ s  $> .1$ ; see Fig. 8, Columns 2–4).

This experiment revealed that the benefit of inflexible and involuntary exogenous attentional orienting remains functionally intact in adults with ADHD. Moreover, this experiment showed for the first time that adults with ADHD possess canonical performance fields (Fig. 9; see [Supplementary Information](#)): Task performance in both groups was better (to an equal extent) at both locations along the horizontal meridian than along the vertical meridian. Both groups were also significantly better at the lower vertical meridian than the upper vertical meridian. Furthermore, the benefit of exogenous attention was similar across locations, thus preserving the shape of the performance fields. These findings are



**Fig. 6** Performance in Experiment 2: Exogenous attention. a Accuracy. b Reaction times. Error bars are  $\pm 1$  SEM. \* $p < .05$ . \*\* $p < .01$ . \*\*\* $p < .001$ . (Color figure online)



**Fig. 7** Valid versus neutral cue condition accuracy for individual observers in Experiment 2. The farther above the diagonal, the greater the attention benefit

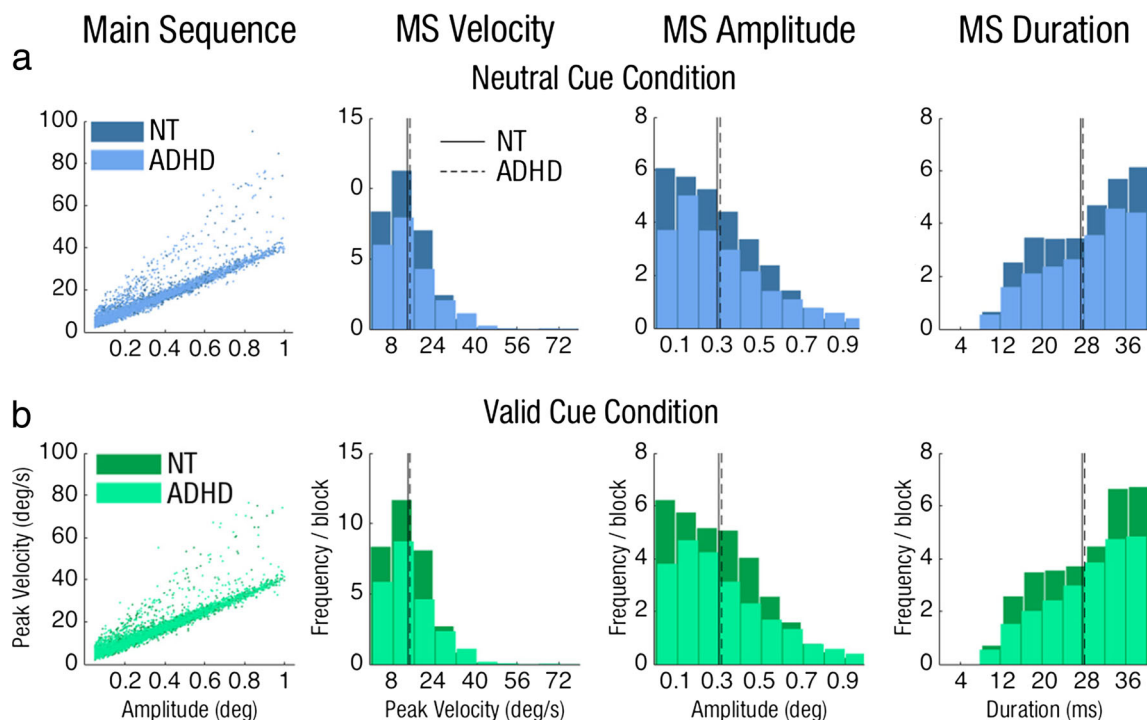
consistent with those of NT adults (Abrams, Nizam, & Carrasco, 2012; Cameron, Tai, & Carrasco, 2002; Carrasco, Giordano, & McElree, 2004; Carrasco et al., 2001).

## General discussion

This study shows that covert orienting and reorienting of *selective* attention—as assessed by our basic task and stimuli—is spared in adults with mild-to-moderate ADHD. This is the first study to

investigate the perceptual effects of both endogenous (Experiment 1) and exogenous (Experiment 2) covert attention in a group of adults with mild-to-moderate ADHD and their age-matched and gender-matched NT controls. We employed a spatial cueing task that is well established in NT observers (Carrasco, 2011) and that has been used to assess selective visual attention in other special populations, i.e., autism spectrum disorder (Grubb, Behrmann, Egan, Minshew, Carrasco, & Heeger, 2013a; b) and individuals with amblyopia (Roberts et al., 2016). The adults with ADHD demonstrated significant and similarly sized benefits of endogenous and exogenous attention to those shown by NT observers. Moreover, in Experiment 1, the cost of deploying endogenous attention to the incorrect target was the same in both groups. A complementary Bayes factor analysis verified that the nonsignificant Group  $\times$  Cue interaction was not due to a lack of statistical power. A recent review of 34 meta-analyses reported that observers with ADHD performed worse than did healthy controls in 96% of the sampled neurocognitive tasks (Pievsky & McGrath, 2017). When weighted by the number of aggregated studies in each meta-analysis, the standardized mean difference (SMD) was .56, a medium-sized effect according to typical benchmarks (Cohen, 1988). Importantly, the mean SMD of the 84 studies with 50 or fewer observers was not meaningfully affected by a smaller sample size.

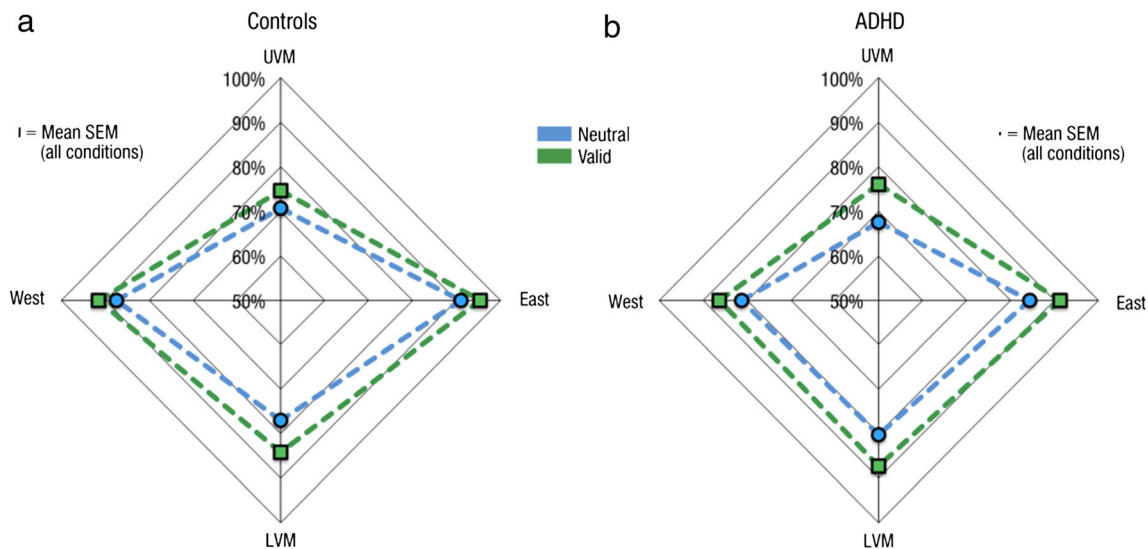
Given reports of substantial differences between ADHD and NT groups on a diverse array of neuropsychological tasks, including some purported to tax “sustained” and “focused”



**Fig. 8** MS kinematics for both groups in Experiment 2: Exogenous attention for the neutral cue condition (a) and valid cue condition (b). MS follow the main sequence in all conditions (Column 1). Distributions of MS velocity (Column 2), amplitude (Column 3), and

duration (Column 4) in terms of average frequency per block for each group. Due to a hardware failure, the data for 13 of 14 observers per group is shown





**Fig. 9** Performance fields for Experiment 2. Accuracy in the valid and neutral cue conditions plotted as a function of target location for the NT controls (a) and adults with ADHD (b)

attention, in studies with similar sample sizes to ours (for reviews, see Frazier et al., 2004; Schoechlin & Engel, 2005), and the results of our Bayesian analyses, we feel confident that our study was powerful enough to detect significant differences between groups, had there been any. In both experiments, we ruled out speed–accuracy trade-offs, and the RT benefits and costs were similar for both groups. Moreover, the overall endogenous RT effect increased with age for both groups. In contrast to reports of greater intraindividual variability in ADHD than controls (Kofler et al., 2013), RT variability did not differ for the two groups. This likely reflects our emphasis on accuracy and the timing of the response window. In summary, voluntary orienting and reorienting, as well as involuntary orienting of covert, selective visual attention, remain functionally intact in adults with ADHD.

Some studies had suggested that orienting and reorienting may be functionally spared in ADHD, but suboptimal methodology limits their interpretation. Three studies on covert attention employed adaptations of the classic Posner spatial cueing task (Posner, 1980), in which an observer must detect a peripheral stimulus as fast as possible while his or her attention is voluntarily or involuntarily drawn to one hemifield via presentation of a spatial cue (Epstein et al., 1997; Epstein et al., 2001; Tomporowski et al., 1994). These studies employed reaction time (RT) as their primary dependent measure; however, RT differences may reflect criterion shifts (Carrasco & McElree, 2001; Wickelgren, 1977), which are more likely in detection than discrimination tasks, and differences in processing speed or sensitivity. Therefore, the reported cueing effects could be attributed to criterion differences between attention conditions rather than to perceptual enhancements. Moreover, the long stimulus-onset asynchronies used (>200 ms) allowed for eye movements, thus potentially confounding the effects attributed to covert attention. Lastly, peripheral cues have been used to study involuntary

attention with a cue-to-stimulus asynchrony (Oberlin et al., 2005) past its maximal effect. Voluntary and involuntary orienting peak by around 300 and 120 ms after cue onset, respectively (Liu et al., 2007; Müller & Rabbitt, 1989; Nakayama & Mackeben, 1989; see a review by Carrasco, 2011). In this study, we overcame all of these methodological concerns.

Only a few studies have directly investigated eye movements in adults with ADHD (e.g., Dankner et al., 2017; Fried et al., 2014; Gooding & Basso, 2008). Some studies have shown that when instructed to move their eyes, individuals in the ADHD group do as well as the NT controls. However, they show some deficits of control with delayed saccade and antisaccade tasks (Gooding & Basso, 2008). In this study, the frequency of saccades did not differ between the groups; this may not be surprising, given the relatively fast temporal demands of our task. Recent studies employing a long sustained attention task found differences in MS rate between ADHD and NT (Danker et al., 2017; Fried et al., 2014). We also analyzed observers' MS to investigate whether these oculomotor correlates of perception would parallel our behavioral results. We found no differences in MS frequency or kinematics between the group of adults with ADHD and NT.

Experiment 2 revealed that in both the neutral and attention conditions, adults with ADHD exhibit the same canonical performance fields as the control group and other NT observers. These novel results indicate that the perceptual sensitivity of ADHD adults as well as the extent and the distribution of exogenous attention are consistent with those of the general population (Abrams et al., 2012; Cameron et al., 2002; Carrasco et al., 2001; Carrasco et al., 2004). In line with findings of preserved perceptual abilities (Kim, Al-Haj, Chen, et al., 2014) and exogenous attention on color appearance (Kim, Al-Haj, Fuller, et al., 2014), the present study rules out an early perceptual or attentional deficit as a contributing factor to explain the diverse symptomatology of the disorder.

Given documented differences in sustained attention and temporal expectation (Dankner et al., 2017; Fried et al., 2014; Hervey et al., 2004; Marchetta et al., 2007; Mueller et al., 2017), as well as literature documenting significant impairments in a diverse set of neurocognitive tasks (for reviews, see Mueller et al., 2017; Pievsky & McGrath, 2017), it is conceivable that differences between groups could emerge with selective attention using harder tasks and/or with more distractors or with an ADHD group with more severe symptomology. Our goal was to isolate the effects of selective attention in ADHD without taxing executive function, and thus these possibilities were outside the scope this study.

A main feature of ADHD, and partly why its diagnosis remains controversial despite decades of research, is that it is heterogeneous (Castellanos & Tannock, 2002; Mueller et al., 2017). Patients diagnosed with the same ADHD label under the current DSM-5 (American Psychiatric Association, 2013) likely suffer from distinct disorders (Milich, Balentine, & Lynam, 2001) with unique severity of symptomology, etiologies, and biological bases; however, there is not a consensus as to the ecological validity of proposed subdivisions (Castellanos & Tannock, 2002). Nevertheless, potential individual differences in selective attention according to subtype and/or severity are open and interesting research questions. The observers for which we had severity scores would mainly be classified as exhibiting mild-to-moderate symptomology (Experiment 2). We cannot report the severity makeup of our ADHD group in Experiment 1, because we could only obtain severity scores for some observers.

We are agnostic regarding whether the underlying neural mechanisms or substrates of attention are the same in adults with ADHD. In fact, substantial anatomical, neuroimaging, and neurophysiological evidence suggest that they are not (Cortese et al., 2012; Mueller et al., 2017). Structurally, studies have found global reductions in gray matter, local gray matter reductions of the prefrontal cortex, anterior cingulate cortex (Seidman et al., 2006), and bilateral early visual areas (Ahrendts et al., 2011), and differences in white matter microarchitecture (Yoncheva et al., 2016). Further, they exhibit abnormal brain activity and disrupted functional connectivity (Konrad & Eickhoff, 2010) between several areas implicated in attentional processing (Cortese et al., 2012). For example, studies have found hypoactivation in dorsolateral and ventral prefrontal cortex, anterior cingulate cortex, and the basal ganglia, as well as hyperactivation in posterior regions of parietal and occipital cortex (Cortese et al., 2012). Further research is needed to link the evidence of difference in brain structure with behavioral differences and similarities between adults with ADHD and NTs.

## Conclusions

The current diagnostic criteria of adult ADHD rely on a combination of both cognitive and neurobehavioral symptoms and

are often assessed using self-report questionnaires and clinical interviews, which can be incomplete, unreliable, and vulnerable to biases. Thus, discovery and clinical implementation of more objective psychometric measures of attentional processes would be valuable. The present study indicates that the perceptual effects of endogenous and exogenous attention are intact in adults with mild-to-moderate ADHD—they improve perception across the visual field. The basic psychophysical attention task we used, if corroborated with larger samples, including observers with more pronounced ADHD symptoms, and tested with other clinical groups, could have translational potential; if incorporated into the clinical diagnostic battery of tests, together with others in which executive function is taxed (e.g., working memory), then it could help in the differential diagnosis of ADHD and other conditions—for example, depression (Hammar & Årdal, 2009; Paelecke-Habermann, Pohl, & Lepow, 2005), schizophrenia (Wang et al., 2005), and anxiety (Pacheco-Unguetti, Acosta, Callejas, & Lupiáñez, 2010)—in which selective attention, although not always optimally manipulated, has been reported to be compromised.

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**Contributions** M.R., B.A., and M.C. designed the study. M.R. and B.A. recruited observers and conducted data collection and analyses. F.X.C. verified ADHD diagnoses. M.R. and M.C. wrote the manuscript; B.A. and F.X.C. provided feedback.

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## SUPPLEMENTARY INFORMATION

### General analysis

We measured accuracy (percentage correct) as the primary dependent variable and average median reaction time (RT), measured relative to the display onset for correct trials only, as the secondary dependent variable. We calculated accuracy and RT for each target location within all three cue conditions, separately for each main experimental block. We then calculated the average accuracy and median RT for each experimental condition as well the mean stimulus contrast (one value per block) across all blocks. Finally, averages for either observer group were calculated separately. All measures of variance are expressed in terms of standard error of the mean (*SEM*) unless specified. In addition to our behavioral measures, we analyzed observers' microsaccades (MS) as they are thought to be informative oculomotor correlates to perception (Dankner, Shalev, Carrasco, & Yuval-Greenberg, 2017; Fried et al., 2014; Gooding & Basso, 2008; Rucci & Poletti, 2015). We defined MS as saccadic eye movements  $>.05^\circ$  but  $<1^\circ$  in amplitude,  $>8^\circ/\text{s}$  but less than  $<40^\circ/\text{s}$  in duration, and with a peak velocity  $<100^\circ$ . Due to a hardware failure, we lost eye-tracking files for some blocks; thus, to compare across individuals and groups, we normalized by the number of analyzable blocks per observer.

## EXPERIMENT 1: ENDOGENOUS ATTENTION

### RT results

Although only a secondary variable, we performed a parallel analysis with RT to rule out speed–accuracy trade-offs. Overall reaction times in the neutral condition were similar between groups. A two-way mixed-design ANOVA revealed a significant main effect of cue,  $F(1.18, 30.7) = 24.7$ ,  $p < .001$ ,  $\eta_p^2 = .49$ , with faster RT in the valid compared to neutral cue condition, and slowest RT for the invalid cue condition, verifying that there were no speed–accuracy trade-offs. Neither the main effect of group ( $F < 1$ ,  $\eta_p^2 = .03$ ) nor the Cue  $\times$  Group interaction,  $F(2, 52) = 0.39$ ,  $p > .1$ ,

$\eta_p^2 = .02$ , was significant (see Fig. 2b in the main text). We found a significant correlation between observers' ages and the overall RT attention effect (invalid – valid cue condition RT) for both the ADHD group,  $r(12) = .61$ ,  $p = .02$ , and the NT control group,  $r(12) = .57$ ,  $p = .03$ ;  $r_s(12) = .61$ ,  $p = .02$ . The older observers were, the more pronounced their RT attention effect. To assess whether intraindividual RT variability differed between the two groups, as has been found in adults with ADHD (Kofler et al., 2013), we calculated the standard deviation of correct trial RTs for each observer. A two-sample  $t$  test found that it did not significantly differ between the observers with ADHD (mean  $SD$  of RT =  $242 \pm 25.9$  ms) and NT controls (mean  $SD$  of RT =  $196 \pm 28.7$ ),  $t(26) = 1.2$ ,  $p > .1$ .

## EXPERIMENT 2: EXOGENOUS ATTENTION

### RT results

Overall RTs in the neutral condition were similar between groups. A two-way mixed-design ANOVA revealed a significant main effect of cue,  $F(1, 26) = 18.9$ ,  $p < .001$ ,  $\eta_p^2 = .42$ , but neither the main effect of group ( $F < 1$ ,  $\eta_p^2 = .02$ ) nor the interaction,  $F(1, 26) = 1.33$ ,  $p > .1$ ,  $\eta_p^2 = .05$ , was significant (see Fig. 6b in the main text). A two-sample  $t$  test showed that the degree of intraindividual RT variability did not significantly differ between the observers with ADHD (mean  $SD$  of RT =  $278 \pm 30.4$  ms) and NT controls (mean  $SD$  of RT =  $227 \pm 34.4$  ms),  $t(26) = 1.1$ ,  $p > .1$ , Cohen's  $d = 0.4$ .

### Investigating potential perceptual and attentional asymmetries “around” the visual field

NT adults exhibit reliable and pronounced differences in accuracy and speed of processing at isoecentric locations as a function of polar angle. These stereotypical “performance fields” (PFs; Abrams, Nizam, & Carrasco, 2012; Carrasco, Giordano, & McElree, 2004; Carrasco, Talgar, & Cameron, 2001), and their underlying neural correlates (Liu, Heeger, & Carrasco, 2006), include both a horizontal-vertical anisotropy (HVA: better performance along the

horizontal than the vertical meridian) and vertical meridian asymmetry (VMA: better performance in the lower than upper visual vertical meridian). In NT individuals, exogenous attention improves accuracy to a similar extent across isoeccentric locations, thereby preserving the overall shape of the PFs (Carrasco et al., 2001; Cameron, Tai, & Carrasco, 2002). Here, we assessed how perceptual sensitivity may differ *around* the visual field as a function of polar angle at isoeccentric locations and whether the magnitude of exogenous attention effects differs across the visual field in adults with ADHD. Comparing their results to those of NT observers will inform us as to whether attention modulates processing to the same degree at locations with different visual sensitivity. To address these questions, we moved the stimuli to the cardinal axes. To maximize statistical power for each of the four target locations, we only tested the valid and neutral cueing conditions.

### **Effect of target location**

A three-way mixed-design ANOVA (Cue × Location × Group) of accuracy revealed significant main effects of cue,  $F(1, 26) = 31.2, p < .001, \eta_p^2 = .55$ , and location,  $F(3, 78) = 22.9, p < .001, \eta_p^2 = .47$ ; neither the main effect of group,  $F(1, 26) = 1.1, p > .1, \eta_p^2 = .04$ , nor any of the two-way or three-way interactions were significant (all  $ps > .1$ , all  $\eta_p^2 < .10$ ; see Fig. 9 in the main text). Bayes factor analysis (Masson, 2011) of the three-way interaction revealed positive evidence in favor of the null hypothesis, that is, a nonsignificant interaction, with odds of 9.9 to 1:  $pBIC(H0|D) = .91$  and  $pBIC(H1|D) = .09$ . A corresponding three-way mixed ANOVA of RT found a similar pattern of results. Thus, the magnitude of the accuracy and RT benefits was similar across locations, and comparable between both groups.

### **Performance fields**

To directly evaluate whether ADHD adults possess canonical PF, that is, better performance at the horizontal than vertical meridian (horizontal-vertical anisotropy, HVA) and at the lower than

upper region of the vertical meridian (vertical meridian asymmetry, VMA), we tested whether both asymmetries significantly differed from zero (one-sample  $t$  tests). Collapsing across cue condition, we averaged performance at the right horizontal meridian (RHM) and left horizontal meridian (LHM), as well as upper vertical meridian (UVM) and lower vertical meridian (LVM), to calculate performance along the horizontal and vertical meridians, respectively. To calculate each observer's HVA, we subtracted his or her average performance along the vertical meridian from the horizontal meridian. To calculate their VMA, we subtracted their performance at the UVM from the LVM.

As expected, the NT controls demonstrated a significant HVA,  $t(13) = 4.44$ ,  $p < .001$ , 95% CI [8.0, 22], Cohen's  $d = 1.2$ , and VMA,  $t(13) = 3.47$ ,  $p = .004$ , 95% CI [3.0, 13], Cohen's  $d = 0.9$  (see Fig. 9a in the main text) in accuracy. The observers with ADHD also exhibited a significant HVA,  $t(13) = 2.97$ ,  $p = .01$ , 95% CI [7.5, 22], Cohen's  $d = 0.8$ , and VMA,  $t(13) = 3.84$ ,  $p = .002$ , 95% CI [5.3, 19], Cohen's  $d = 1.0$  (see Fig. 9b in the main text) in accuracy. Again, there were no speed–accuracy trade-offs; for RTs, the NT controls had neither a significant HVA,  $t(13) = 1.76$ ,  $p > .1$ , or VMA,  $t(13) = 1.53$ ,  $p > .1$ ; the ADHD observers exhibited a significant HVA,  $t(13) = 3.16$ ,  $p = .008$ , and VMA,  $t(13) = 2.27$ ,  $p = .04$ , with faster times at higher accuracy locations. Moreover, the benefit of attention did not differ as a function of target location within either group (Cue  $\times$  Location), NT,  $F < 1$ ; ADHD,  $F(3, 39) = 1.0$ ,  $p > .1$ .



**Table S1.** Clinical and demographic information for adult observers with ADHD

Observer <sup>1</sup>	Gender	Age	Race/Ethnicity <sup>2</sup>	Handedness <sup>3</sup>	Subtype <sup>4</sup>	Medication <sup>5</sup>	Comorbidities <sup>6</sup>
Participated in both experiments							
1	M	22	C / NH	L	C	SM	MDD, GAD, SA
2	F	29	C / NH	R	I	NN	MDD, GAD
3	F	26	C / NH	R	I	NN	NR
4	F	53	C / NH	R	I	NN	None
5	M	39	C / NH	R	I	MN	None
6	F	21	M / NH	R	I	SM	None
Total	4 F, 2 M	$M = 31.7,$ $SD = 12.3$	5 C, 1 M / 6 NH	5 R, 1 L	5 I, 0 H, 1 C, 0 N	1 MN, 3 NN, 2 SM	2 Comorbidities, 3 None, 1 NR
Participated only in Experiment 1							
7	M	31	C, NH	R	I	NN	None
8	M	40	M, NH	R	H	MN	None
9	M	31	C, NH	R	N	SM	PTSD
10	M	31	C, NH	R	I	SM	None
11	M	24	C, NH	R	I	SM	None
12	F	28	C, NH	R	I	MN	None
13	F	25	C, NH	R	I	SM	TD
14	F	27	A, NH	R	C	MN	None
Total in Exp 1	7 F, 7 M	$M = 30.5,$ $SD = 8.5$	11 C, 2 M, 1 A / 14 NH	13 R, 1 L	10 I, 1 H, 2 C, 1 N	4 MN, 4 NN, 6 SM	4 Comorbidities, 9 None, 1 NR
Participated only in Experiment 2							
7	M	40	C / NH	R	C	SM	MDE
8	M	24	NR / H	NR	I	NN	MDE, AD
9	M	29	C / NH	R	I	SM	None
10	F	29	NR / NR	NR	N	SM	rADJ, MDE
11	F	29	NR / NR	NR	C	NN	None
12	M	36	C / NH	R	I	SM	None
13	F	37	C / NH	R	I	SM	GAD, PA, CD, OCD
14	M	25	C / NH	L	I	MN	rMDD, rANX
Total in Exp 2	7 F, 7 M	$M = 31.4,$ $SD = 8.7$	10 C, 1 M, 3 NR / 10 NH, 1 H, 3 NR	9 R, 2 L, 3 NR	10 I, 0 H, 3 C, 1 N	2 MN, 5 NN, 7 SM	7 Comorbidities, 6 None, 1 NR

<sup>1</sup>Observers were clinically diagnosed with *DSM-IV-TR* ADHD (APA, 2000) according to the Adult ADHD Clinical Diagnostic Scale v.1.2 (Adler & Spencer, 2004), and Structured Clinical Interview for DSM-IV, Research Version, Non-patient Edition (SCID-I/NP; First et al., 2002).

<sup>2</sup>Race: A = Asian, C = Caucasian, M = Mixed Race, NR = Not Reported; Ethnicity: H = Hispanic, NH = Not Hispanic, NR = Not Reported

<sup>3</sup>R = Right-handed, L = Left-handed, NR = Not Reported

<sup>4</sup>I = Predominantly Inattentive, H = Predominantly Hyperactive/Impulsive, C = Combined

<sup>5</sup>MN = Medication Naïve, NN = Not Naïve But Off Medication At Time Of Experiment, SM = Prescribed Stimulant Medication  
Patients abstained from medication within 24 hrs of the exp.

<sup>6</sup>r = In Remission; ADJ = Adjustment Disorder, AF = Affective Disorder, ANX = Anorexia Nervosa, CD = Childhood Depression, GAD = Generalized Anxiety Disorder, MDD = Major Depressive Disorder, MDE = Experienced  $\geq 1$  Major Depressive Episodes, OCD = Obsessive Compulsive Disorder, PD = Panic Disorder, PTDS = Post-Traumatic Stress Disorder, SA = Substance Abuse, TD = Trauma Disorder, Otherwise Unspecified

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